Search for Differences in Oscillation Parameters for Atmospheric Neutrinos and Antineutrinos at Super-Kamiokande

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We present a search for differences in the oscillations of antineutrinos and neutrinos in the Super-Kamiokande I, II, and III atmospheric neutrino sample. Under a two-flavor disappearance model with separate mixing parameters between neutrinos and antineutrinos, we find no evidence for a difference in oscillation parameters. Best fit antineutrino mixing is found to be at \((\Delta m^2, \sin^2 \theta) = (2.0 \times 10^{-3} \text{eV}^2, 1.0)\) and is consistent with the overall Super-K measurement.

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As the parameters outlining the standard neutrino oscillation framework become increasingly well known, searches for sub-leading and possibly symmetry-breaking effects become possible. If the value of \( \theta_{13} \) is non-zero, for instance, it becomes possible to search for CP-violation effects in the neutrino system via differences in the oscillation probabilities of neutrinos and antineutrinos. In this paper we consider the possibility that the survival probability \( P(\nu_\mu \rightarrow \nu_\mu) \) is governed by a different mass splitting or mixing angle compared to \( P(\bar{\nu}_\mu \rightarrow \bar{\nu}_\mu) \). This is not considered in most oscillation studies as the mass splitting and mixing angle are expected to be identical for neutrinos and antineutrinos by CPT symmetry. An inequality of these probabilities, in the absence of matter effects, could signal new physics. For atmospheric muon neutrino disappearance, which is predominantly oscillation into tau neutrinos, the matter effect is expected to be small and the mixing parameters for muon neutrino and antineutrino disappearance appearance should be the same. The MINOS experiment, which is sensitive to neutrino oscillations at the atmospheric scale, and which can determine the sign of muons by magnetic bending, has observed antineutrino disappearance [1] at a best-fit value of \( \Delta m^2 \) nearly 50\% larger than previous measurements made using neutrinos [2–5]. Though not in the realm of atmospheric mixing, the MiniBooNE experiment has similarly observed a discrepancy between its neutrino [6] and antineutrino [7] data. Therefore, further tests of differences between neutrinos and antineutrinos using atmospheric data are well motivated.

Super-Kamiokande (Super-K, SK), described below, cannot distinguish \( \nu \) from \( \bar{\nu} \) on an event by event basis so potential differences in their oscillations would appear in the atmospheric neutrino sample in a statistical way. Notably, the neutrino and antineutrino cross sections differ by a factor of two to three depending on the neutrino energy. The ratio of the \( \nu \) and \( \bar{\nu} \) atmospheric fluxes is similarly energy dependent [8]. For these reasons, even in the absence of CPT-violating oscillations the relative numbers of each species are expected to differ. Kinematic considerations can also induce differences in the products of neutrinos and antineutrinos, enhancing the sample purity of one or the other. For instance, the absorption of \( \pi^- \) on \(^{16}\)O nuclei in water tends to enrich the neutrino component of samples that are sub-divided based on their number of decay electrons. If CPT-violating oscillations are present in the data, the distortion of the zenith angle distribution characteristic of \( \nu_\mu \) disappearance would appear at different energies and path lengths (different oscillation frequencies) between neutrinos and antineutrinos. Since Super-K can only observe the total distribution, a potential signal would appear as a distortion consistent with the composition of separately oscillated spectra. In this context, antineutrinos, which have the smaller cross section, are expected to provide a weaker oscillation constraint than neutrinos.

In this Letter we consider \textit{ad-hoc} CPT-violating oscillations by testing separate two-neutrino disappearance models for neutrinos and antineutrinos:

\[
P(\nu_\mu \rightarrow \nu_\mu) = 1 - \sin^2 2\theta \sin \left( \frac{\Delta m^2 L}{4E} \right)
\]

\[
P(\bar{\nu}_\mu \rightarrow \bar{\nu}_\mu) = 1 - \sin^2 2\bar{\theta} \sin \left( \frac{\Delta \bar{m}^2 L}{4E} \right),
\]

where \( L \) is the neutrino path-length and \( E \) is the neutrino energy. In the presence of matter, additional neutrino-electron scattering induces a CPT-violating-like difference between the neutrino and antineutrino survival probabilities [9, 10], particularly when \( \theta_{13} \) is non-zero. Recent data suggest [11–13] that \( \theta_{13} \) is small and therefore matter-induced corrections to Eqn. 1 are expected to be sub-dominant. Changes to the fit results induced by a three-flavor treatment are briefly considered below. Although the presence of both matter and solar mixing terms is expected to drive \( \nu_\mu \rightarrow \nu_e \) transitions below about 1 GeV even if \( \theta_{13} = 0 \), the oscillation frequency in this domain is high enough that the effects are averaged out by the detector resolution. Since the dominant atmospheric disappearance effect is seen at higher energies, a two-neutrino scheme is the focus of this study.

Super-Kamiokande is a water Cherenkov detector located in Japan’s Gifu prefecture and situated at a depth of 2700 meters water equivalent. It is comprised of two concentric, optically separated cylinders: an inner detector (ID) viewing a 22.5 kton fiducial volume and an outer detector (OD) used primarily as a veto. During the first run of the detector, SK-I, the walls of the ID were lined with 11,146 inward-facing 20-inch photomultiplier tubes (PMTs). The two subsequent run periods, SK-II and SK-III, had 5,182 and 11,129 ID PMTs, respectively and the PMTs have been encased in fiber-reinforced plastic shells. The OD has been instrumented with 1,885 outward-facing 8-inch PMTs throughout. More details on the detector and its calibration may be found in [14].

The atmospheric neutrino data are divided into three categories. Fully contained (FC) events deposit all of their light in the ID, partially contained (PC) events correspond to 1489 live-days with 1646 days of exposure, and through-going (Up) events are required to have a minimum path-length of seven meters and are classified as stopping or through-going. Through-going Up events are further sub-divided into “showering” and “non-showering” based on [15].

The present analysis uses data accumulated during the first three SK run periods. SK-I (1996 to 2001) FC and PC events correspond to 1489 live-days with 1646 days of Up events. SK-II (2002 to 2005) had 799 FC/PC and 828 live-days of Up events. The FC and PC lifetimes
FIG. 1: SK-I+II+III lepton momentum (first column) and zenith angle distributions of the event samples used in the analysis. Black dots represent the data with statistical errors, the histogram is the oscillated MC expectation at the best fit point with the shaded region showing the antineutrino composition.

during SK-III (2005-2007) were 518 days and that for Up-µ was 636 days.

During the analysis, FC and PC events are further divided. Fully contained events are separated into sub-GeV \((E_{vis} < 1.33 \text{ GeV})\) and multi-GeV \((E_{vis} > 1.33 \text{ GeV})\). These samples are then separated based on their number of reconstructed Cherenkov rings into single- and multi-ring topologies. Pattern-identification of single-ring events is used to separate them into e-like and µ-like categories. This technique is applied to the multi-ring sample using the most energetic Cherenkov ring. The sub-GeV single-ring e-like and µ-like samples are also divided based upon their number of decay-electrons. A π⁰-like sample is also extracted from the single-ring e-like events [12]. Partially contained events are separated into “OD stopping” and “OD through-going” categories based on the amount of Cherenkov light observed in the OD at the exit point.

Since the physical configuration of the detector and its reconstruction performance varies among the SK run periods, separate 500 year MC samples are used for each. A “pulled” \(\chi^2\) [16] based on a Poisson probability distribution is used to compare the data against the MC:

\[
\chi^2 = 2 \sum_n \left( \sum_i E_{n}^{SKi}(1 + \sum_j f^n_i \epsilon_j) - \sum_i O_{n}^{SKi} \ln \frac{\sum_i O_{n}^{SKi}}{\sum_i E_{n}^{SKi}(1 + \sum_j f^n_i \epsilon_j)} \right) + \sum_k \left( \frac{e_k}{\sigma_k} \right)^2 . \tag{2}
\]

In this equation \(n\) indexes the data bins, \(E_{n}^{SKi}\) is the MC expectation for SK-\(i\), and \(O_{n}^{SKi}\) is the number of observed
events in the \( n^{th} \) bin during SK-1. The effect of the \( i^{th} \) systematic error is introduced via the error parameter \( \epsilon_i \) and \( f^i_n \), where the latter is the fractional change in the MC expectation of bin \( n \) introduced by a 1-sigma shift in its systematic error, \( \sigma_i \). The data and MC are divided into 420 bins for each of the SK run periods when computing these systematic errors, but are later merged as above to ensure the stability of the fit function against sparsely populated bins. In total, 420 bins are used to compute the value of \( \chi^2 \).

Equation (2) is minimized with respect to the \( \epsilon_i \) according to \( \frac{\partial \chi^2}{\partial \epsilon_i} = 0 \), yielding a set of linear equations in \( \epsilon_i \) that are solved iteratively [16]. Following this procedure a \( \chi^2 \) value is computed for each point in the oscillation parameter space. The global minimum \( \chi^2 \) is defined as the analysis’ best fit point.

<table>
<thead>
<tr>
<th>Parameter Best Fit</th>
<th>90% C.L.</th>
<th>Three-Flavor</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \Delta m^2(eV^2) )</td>
<td>( 2.1 \times 10^{-3} \times [1.7, 3.0] \times 10^{-3} \times [1.7, 3.3] \times 10^{-3} )</td>
<td></td>
</tr>
<tr>
<td>( \Delta m^2(eV^2) )</td>
<td>( 2.0 \times 10^{-3} \times [1.3, 4.0] \times 10^{-3} \times [1.2, 4.0] \times 10^{-3} )</td>
<td></td>
</tr>
<tr>
<td>( \sin^2 \theta_1 )</td>
<td>1.0 \times [0.93, 1.0] \times [0.93, 1.0]</td>
<td></td>
</tr>
<tr>
<td>( \sin^2 \theta_2 )</td>
<td>1.0 \times [0.83, 1.0] \times [0.78, 1.0]</td>
<td></td>
</tr>
</tbody>
</table>

TABLE I: Best fit information for the four parameter fit to the SK-I+II+III data set. The 68\%, 90\%, and 99\% allowed region appear in thin, medium, and thick lines respectively. The shaded region shows the 90\% C.L. allowed region for antineutrino disappearance in an antineutrino beam from MINOS [1]. A solid point denotes the location of the best fit from that analysis.

The 120 sources of systematic uncertainty considered in this analysis are separated into two classes: those that are common throughout the SK run periods and those that are dependent upon a particular detector geometry. Common systematic errors stem from uncertainties in the neutrino interaction cross sections, nuclear effects, and the atmospheric neutrino flux. Independent systematic errors are related to detector performance and include uncertainties in the event reconstruction and reduction. A complete list of the systematic errors used here is presented in [12].

Antineutrino oscillations are considered independently of neutrino oscillations over a four-dimensional oscillation space with two parameters for each: \( (\Delta m^2, \sin^2 \theta) \) and \( (\Delta m^2, \sin^2 \theta) \). All parameters are varied simultaneously on a grid of 50 \times 35 points in the antineutrino plane and 20 \times 10 points in the neutrino plane. The neutrino(antineutrino) parameter space is taken over \( 1.0(0.7) \times 10^{-3} \leq \Delta m^2 \leq 5.0(8.0) \times 10^{-3} eV^2 \) and \( 0.85(0.65) \leq \sin^2 \theta \leq 1.0(1.0) \), comprising an area encompassing the current allowed values of these parameters [5, 12]. Minimizing the \( \chi^2 \) function in Eqn. 2 over this parameter space, the best fit is found with \( \chi^2 = 468.4 \) for 416 degrees of freedom. Table I summarizes the fit information. Since the oscillations of atmospheric neutrinos are sensitive to the effects of \( \theta_{13} \) and \( \delta_{\text{cp}} \), an additional analysis has been performed assuming distinct three-flavor mixing between neutrinos and antineutrinos. Though this model has no strong theoretical motivation, the potential impact of a three-flavor framework on the SK allowed regions is provided in the table for reference. In both fits no difference between antineutrino and neutrino mixing is found in the data.

The combined data overlaid with the MC expectation at the best fit point are shown in Fig. 1. Antineutrinos have been oscillated independently of neutrinos in the expectation and are shown as the shaded portion of the histogram. Figure 2 shows the allowed regions at several C.L. in the antineutrino plane. The neutrino parameters have been minimized over and the contours have been drawn for a standard \( \Delta \chi^2 \) distribution with two degrees of freedom. The 90\% C.L. regions from the antineutrino and neutrino parameters overlaid with the allowed region from the standard two-flavor analysis, where the mixing parameters are required to be identical for neutrinos and antineutrinos, appear in Fig. 3. The best fit point of the standard analysis, \( (\Delta m^2, \sin^2 \theta) = (2.1 \times 10^{-3} eV^2, 1.0) \), is consistent with the best fit from the antineutrino separated analysis.

At each point in the plane of Fig. 2 the best fit to the data may lie at a point in the neutrino parameters that does not correspond to equal neutrino and antineutrino mixing. To illustrate the difference between neutrino and antineutrino oscillations permitted by the data,
In conclusion, a search for evidence of differing neutrino and antineutrino oscillation parameters in the SK-I+II+III atmospheric data sample has been carried out. The atmospheric mixing parameters for antineutrino oscillations are consistent with those for neutrino oscillation. The results agree with the standard SK atmospheric oscillation analysis, which have also been presented. The SK antineutrino oscillation best fit is consistent with the parameters found by MINOS using a predominantly muon neutrino beam.

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