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# Deuterium target data for precision neutrino-nucleus cross sections 

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#### Abstract

Amplitudes derived from scattering data on elementary targets are basic inputs to neutrinonucleus cross section predictions. A prominent example is the isovector axial nucleon form factor, $F_{A}\left(q^{2}\right)$, which controls charged current signal processes at accelerator-based neutrino oscillation experiments. Previous extractions of $F_{A}$ from neutrino-deuteron scattering data rely on a dipole shape assumption that introduces an unquantified error. A new analysis of world data for neutrinodeuteron scattering is performed using a model-independent, and systematically improvable, representation of $F_{A}$. A complete error budget for the nucleon isovector axial radius leads to $r_{A}^{2}=$ $0.46(22) \mathrm{fm}^{2}$, with a much larger uncertainty than determined in the original analyses. The quasielastic neutrino-neutron cross section is determined as $\left.\sigma\left(\nu_{\mu} n \rightarrow \mu^{-} p\right)\right|_{E_{\nu}=1 \mathrm{GeV}}=10.1(0.9) \times 10^{-39} \mathrm{~cm}^{2}$. The propagation of nucleon-level constraints and uncertainties to nuclear cross sections is illustrated using MINERvA data and the GENIE event generator. These techniques can be readily extended to other amplitudes and processes.


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## I. INTRODUCTION

Current and next generation accelerator-based neutrino experiments are poised to answer fundamental questions about neutrinos [1-5]. Precise neutrino scattering cross sections on target nuclei are critical to the success of these experiments. These cross sections are computed using nucleon-level amplitudes combined with nuclear models. Determination of the requisite nuclear corrections presently relies on data-driven modeling [6-9] employing experimental constraints [10-17]. Ab-initio nuclear computations are beginning to provide additional insight [1820]. Regardless of whether nuclear corrections are constrained experimentally or derived from first principles, independent knowledge of the elementary nucleon-level amplitudes is essential. In this paper, we address the problem of model-independent extraction of elementary amplitudes from scattering data, and the propagation of rigorous uncertainties through to nuclear observables.

[^0]The axial-vector nucleon form factor, $F_{A}\left(q^{2}\right)$, is a prominent source of uncertainty in any neutrino cross section program. While the techniques employed in the present paper may be similarly applied to other elementary amplitudes, such as vector form factors [21], we focus on the axial-vector form factor, which is not probed directly in electron scattering measurements, and which has large uncertainty.

The axial form factor is constrained, with a varying degree of model dependence, by neutron beta decay [22], neutrino scattering on nuclear targets heavier than deuterium [11, 23-28], pion electroproduction [29] and muon capture [30]. Existing data for the neutrinodeuteron scattering process provide the most direct access to the shape of the axial-vector nucleon form factor. The assumption of a neutron at rest and barely bound in the laboratory frame permits unambiguous energy reconstruction, eliminating flux uncertainties. The abundant neutrino scattering data on heavier targets involve degenerate uncertainties from neutrino flux, and from large and model-dependent nuclear corrections, complicating the extraction of nucleon-level amplitudes. Antineutrino scattering on hydrogen would entirely eliminate even the nuclear corrections required for deuterium, but there are no high-statistics data for this process. Given the im-
portance of deuterium data for the axial form factor, it is imperative to quantify the constraints from existing data.

In this paper, we present the charged-current axialvector nucleon form factor and error budget determined from neutrino-deuterium scattering data. In place of the dipole assumption (cf. Eq. (9) below) used in previous analyses of the form factor, we employ the modelindependent $z$ expansion ${ }^{1}$ parameterization. The resulting uncertainty is significantly larger than found in previous analyses [29, 54, 55] of the deuterium constraint on the axial form factor using multiple data sets. This larger uncertainty results from removing the dipole assumption, and from including systematic errors for experimental acceptance corrections and for model-dependent deuteron corrections. The new constraints may be readily implemented in nuclear models and neutrino event generators.

The remainder of the paper is structured as follows. In Sec. II we introduce the deuterium datasets and perform fits to the dipole model for the axial form factor. This is done in order to compare with original publications, and to isolate the impact of form factor shape assumptions versus other inputs or data selections. In Sec. III we review the relevant $z$ expansion formalism, and redo fits from Sec. II replacing dipole with $z$ expansion. Several features of these fits indicate potentially underestimated systematic errors in corrections that were applied to data in the original publications. Section IV describes a range of systematics tests. We consider several sources of systematic errors in more detail in Sec. V, and redo fits in Sec. VI, where we present final results for $F_{A}\left(q^{2}\right)$. In Sec. VII we illustrate the propagation of errors to several derived observables, including the isovector axial nucleon radius and total neutrino-nucleon quasielastic cross sections. The incorporation of nucleon-level uncertainties in nuclear cross sections is illustrated with MINERvA data [56]. Section VIII provides a summary and conclusion.

## II. DEUTERIUM DATA AND DIPOLE FITS

The world data from deuterium bubble chamber experiments consists of deuterium fills of the ANL 12foot deuterium bubble chamber experiment [57-59], the BNL 7-foot deuterium bubble chamber experiment [60], and the FNAL 15 -foot deuterium bubble chamber experiment [61]. We refer below to these experiments as ANL1982, BNL1981 and FNAL1983, respectively. ${ }^{2}$

[^1]
## A. Fits to $Q^{2}$ distributions

Extracting the axial form factor from data requires information about all other aspects of the scattering crosssection. The original publications used a variety of different inputs for axial $\left(g_{A}\right)$ and magnetic $\left(\mu_{p}-\mu_{n}\right)$ couplings, vector and pseudoscalar form factors, nuclear corrections, and muon mass corrections. Table I displays the input choices made in the original publications for each of the three considered datasets, as well as the updated inputs used for the remainder of this paper. ${ }^{3}$

The vector form factors are constrained by invoking isospin symmetry and constraints of electron-nucleon scattering data. In place of the Olsson vector form factors [63], we use the so-called BBA2005 parameterization that is commonly employed in contemporary neutrino studies [64]. Similar results were obtained using the BBA2003 [65] and BBBA2007 [66] parameterizations. Recent developments, connected with the socalled "proton radius puzzle", point to potential shortcomings in previous extractions of the vector form factors $[35,67,68]$. A systematic study of the vector form factors similar to the $z$ expansion analysis of the axial form factor presented here is undertaken in Refs. [21, 35].

For the pseudoscalar form factor $F_{P}$, we employ the partially conserved axial current (PCAC) ansatz,

$$
\begin{equation*}
F_{P}^{\mathrm{PCAC}}\left(q^{2}\right)=\frac{2 m_{N}^{2} F_{A}\left(q^{2}\right)}{m_{\pi}^{2}-q^{2}} \tag{1}
\end{equation*}
$$

The free-nucleon form factors $F_{A}$ and $F_{P}$ are functions of the four momentum transfer $q^{2}$ from the lepton to the nucleon, and $m_{N}=0.9389 \mathrm{GeV}, m_{\pi}=0.14 \mathrm{GeV}$ are the masses of the nucleon and the pion. The effects of the pseudoscalar form factor are suppressed in the limit of small lepton mass, and its uncertainties are negligible in most applications involving accelerator neutrino beams, including this analysis.

Nuclear corrections relating the free neutron cross section, $d \sigma^{n}$, to the deuteron cross section, $d \sigma^{D}$, may be parameterized as

$$
\begin{equation*}
\frac{d \sigma^{D}}{d Q^{2}}=R\left(Q^{2}, E_{\nu}\right) \frac{d \sigma^{n}}{d Q^{2}} \tag{2}
\end{equation*}
$$

where $d \sigma^{D} / d Q^{2}$ denotes the deuteron differential cross section with respect to the intrinsically positive $Q^{2}=$ $-q^{2} .{ }^{4}$ The model of Ref. [69] was used in the original
(cf. Ref. [62], Fig. 5). We thus consider only the events from the BNL1981 dataset.
${ }^{3}$ Form factor notations and conventions are as in Ref. [31].
${ }^{4}$ For definiteness in the deuteron case, we let $Q^{2}$ in Eq. (2) denote the leptonic momentum transfer. This definition is consistent with the experimental reconstruction, which assumed the kinematics for scattering from a free neutron in the presence of a spectator proton carrying opposite momentum to the neutron.

TABLE I. Inputs from the original publications, BNL1981 [60], ANL1982 [59] and FNAL1983 [61], and our default inputs. See text for details.

| input | BNL1981 | ANL1982 | FNAL1983 | this work | reference |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $g_{A}=F_{A}(0)$ | -1.23 | -1.23 | -1.23 | -1.2723 | $[22]$ |
| $\mu_{p}-\mu_{n}-1$ | 3.708 | 3.71 | 3.708 | 3.7058 | $[22]$ |
| $F_{V i}$ | Olsson [63] | Olsson [63] | Olsson [63] | BBA2005 | $[64]$ |
| $F_{P}$ | PCAC | PCAC | PCAC | PCAC | $(1)$ |
| Deuteron correction | Singh [69] | Singh [69] | Singh $[69]$ | Singh | $[69]$ |
| lepton mass | $m_{\mu}=m_{\mu}$ except ABC | $m_{\mu}=m_{\mu}$ | $m_{\mu}=m_{\mu}$ except ABC | $m_{\mu}=m_{\mu}$ |  |
| $Q^{2}$ range | $0.06-3 \mathrm{GeV}^{2}$ | $0.05-2.5 \mathrm{GeV}^{2}$ | $0-3 \mathrm{GeV}^{2}$ |  |  |
| $N_{\text {bins }}$ | 49 | 49 | 30 |  |  |
| $N_{\text {events }}$ | 1236 | 1792 | 354 |  |  |
| kinematic cut | $Q^{2} \geq 0.06 \mathrm{GeV}^{2}$ | $Q^{2} \geq 0.05 \mathrm{GeV}^{2}$ | $Q^{2} \geq 0.10 \mathrm{GeV}^{2}$ |  |  |

analyses, with $R\left(Q^{2}, E_{\nu}\right) \approx R\left(Q^{2}\right)$ independent of neutrino energy, and $R\left(Q^{2}\right) \rightarrow 1$ above $Q^{2} \approx 0.2 \mathrm{GeV}^{2}$. We retain this model as default, but examine deviations from this simple description below in Sec. IV, using the calculations of Ref. [70].

The neutrino-neutron quasielastic cross section may be written in a standard form

$$
\begin{equation*}
\frac{d \sigma^{n}}{d Q^{2}} \propto \frac{1}{E_{\nu}^{2}}\left[A\left(Q^{2}\right) \mp B\left(Q^{2}\right) \frac{s-u}{m_{N}^{2}}+C\left(Q^{2}\right) \frac{(s-u)^{2}}{m_{N}^{4}}\right] \tag{3}
\end{equation*}
$$

where $s-u=4 E_{\nu} m_{N}-Q^{2}-m_{\mu}^{2}$ is the difference of Mandelstam variables, $A, B$ and $C$ are quadratic functions of nucleon form factors [71], and the vector-axial interference term $B$ changes sign for the $\bar{\nu} p$ scattering process. In the BNL1981 and FNAL1983 datasets, the lepton mass was neglected inside the functions $A\left(Q^{2}\right)$, $B\left(Q^{2}\right)$ and $C\left(Q^{2}\right)$ of Eq. (3), but retained in other kinematic prefactors. In our analysis, we retain the complete lepton mass dependence.

The event distributions in $Q^{2}$ have been obtained by digitizing the relevant plots from the original publications. Table I gives the $Q^{2}$ range and bin size, the total number of events, ${ }^{5}$ and the minimum $Q^{2}$ retained in the original analyses. In each case, events in a lowest $Q^{2}$ bin were omitted from fits, and only FNAL1983 reports these events. We retain the same binning and minimum $Q^{2}$ cut in our default fits. These distributions are included as supplementary material to the present paper [72].

[^2]
## B. $E_{\nu}$ distributions and flux

An advantage of the $\nu_{\mu} d \rightarrow \mu^{-} p p$ process in an exquisite device like a bubble chamber is the accurate reconstruction of the neutrino energy for each event. Cross section parameters can be constrained from the $Q^{2}$ distribution despite poorly-controlled uncertainties in ab-initio neutrino flux estimates. This is especially valuable for the low energy ANL1982 and BNL1981 data, whose neutrino energy spectrum significantly influences the shape of the $d N / d Q^{2}$ distribution through the energy-dependent kinematic limit corresponding to a backscattered lepton.

Unfortunately, event-level kinematics from the deuterium datasets are no longer available and unbinned likelihood fits using the $E_{\nu}$ and $Q^{2}$ dependence of the cross section cannot be repeated. However, the onedimensional distribution of events in reconstructed neutrino energy, $d N / d E_{\nu}$, may be extracted from the original publications, and we use this information to reconstruct the flux self-consistently. This subsection describes the procedure we use, including some subtle points required for later interpretation of the form factor fits.

The differential neutrino flux is determined by

$$
\begin{equation*}
\frac{d \Phi\left(E_{\nu}\right)}{d E_{\nu}} \propto \frac{1}{\sigma^{n}\left(E_{\nu}, F_{A}\right)} \frac{d N^{n}}{d E_{\nu}} \tag{4}
\end{equation*}
$$

where $\sigma^{n}\left(E_{\nu}, F_{A}\right)$ is the free-neutron quasielastic cross section, and $d N^{n} / d E_{\nu}$ is the energy distribution of freeneutron events that would be obtained in the experimental flux. The constant of proportionality in Eq. (4) is determined by the number of target deuterons and the time duration of the experiment. Let us normalize the energy distribution according to

$$
\begin{equation*}
\int_{0}^{\infty} d E_{\nu} \frac{d N^{n}}{d E_{\nu}}=\mathcal{N} \int_{Q_{\min }^{2}}^{\infty} d Q^{2} \frac{d N^{D}}{d Q^{2}} \tag{5}
\end{equation*}
$$

Consistency in Eq. (5) is obtained when $\mathcal{N}=$
$\hat{\mathcal{N}}\left(F_{A}, Q_{\text {min }}^{2}, R\right)$, where

$$
\begin{equation*}
\hat{\mathcal{N}}\left(F_{A}, Q_{\min }^{2}, R\right)=\frac{\int_{0}^{\infty} d Q^{2} \frac{d N^{n}}{d Q^{2}}}{\int_{Q_{\min }^{2}}^{\infty} d Q^{2} \frac{d N^{D}}{d Q^{2}}} \tag{6}
\end{equation*}
$$

The right hand side of Eq. (6) may be computed using a given $d N^{n} / d E_{\nu}$, and depends on $R\left(Q^{2}, E_{\nu}\right)$ and $F_{A}\left(q^{2}\right)$ through Eqs. (2) and (4). Using the flux from Eq. (4), we have finally,

$$
\begin{align*}
& \left(\frac{d N^{D}}{d Q^{2}}\right)^{\text {theory }} \\
& =\mathcal{N}_{\text {fit }} \int_{0}^{\infty} d E_{\nu} \frac{R\left(Q^{2}, E_{\nu}\right) \frac{d \sigma^{n}}{d Q^{2}}\left(E_{\nu}, F_{A}, Q^{2}\right)}{\sigma^{n}\left(E_{\nu}, F_{A}\right)} \frac{d N^{n}}{d E_{\nu}} \tag{7}
\end{align*}
$$

where a fit parameter, $\mathcal{N}_{\text {fit }}$, has been introduced for the normalization.

Choosing $\mathcal{N}=\hat{\mathcal{N}}\left(F_{A}, Q_{\text {min }}^{2}, R\right)$ in Eq. (5) would correspond to $\mathcal{N}_{\text {fit }}=1$. In order to avoid the explicit computation of the integrals (6), we instead take $\mathcal{N}=1$, corresponding to the expectation $\mathcal{N}_{\text {fit }}=\mathcal{N}\left(F_{A}, Q_{\text {min }}^{2}, R\right)$. We allow the parameter $\mathcal{N}_{\text {fit }}$ to float unconstrained in the fits, with an independent parameter for each experiment.

We emphasize that $d N^{n} / d E_{\nu}$ in Eq. (4) represents the energy distribution of free-neutron events that would be obtained in the experimental flux; this distribution is obtained from the energy distribution of observed events in deuterium by correcting for nuclear effects, for events lost due to the $Q_{\text {min }}^{2}$ cut, and for other experimental effects. Such corrections were applied to the energy distribution presented in the BNL1981 dataset, but not in the ANL1982 and FNAL1983 datasets. The effect of applying or not applying these corrections is found to be small, as discussed below in Sec. V A.

For later comparison, we compute the ratios (6) with a nominal dipole axial form factor $\left(m_{A}=1 \mathrm{GeV}\right.$, cf. Eq. (9) below), neglecting deuteron corrections ( $R=1$ ), and at a nominal $E_{\nu}=1 \mathrm{GeV}$ neutrino energy, for the $Q_{\text {min }}^{2}$ values employed in the BNL1981, ANL1982, FNAL1983 datasets: ${ }^{6}$

$$
\begin{align*}
& \hat{\mathcal{N}}\left(Q_{\text {min }}^{2}=0.06 \mathrm{GeV}^{2}\right) \approx 1.13 \\
& \hat{\mathcal{N}}\left(Q_{\text {min }}^{2}=0.05 \mathrm{GeV}^{2}\right) \approx 1.11 \\
& \hat{\mathcal{N}}\left(Q_{\text {min }}^{2}=0.10 \mathrm{GeV}^{2}\right) \approx 1.23 \tag{8}
\end{align*}
$$

We expect these numbers to be approximately reproduced in $\mathcal{N}_{\text {fit }}$ when the deviation from $\hat{N}=1$ in Eq. (6) is dominated by the $Q_{\text {min }}^{2}$ cut.

[^3]TABLE II. Dipole axial mass extracted in original publications, our extraction using parameter inputs as in the original publications, and our extraction using updated constants and vector form factors as in Table I. Errors are statistical only.

|  | $m_{A}^{\text {dipole }}($ ref $)$ | $m_{A}^{\text {dipole }}($ old $)$ | $m_{A}^{\text {dipole }}($ new $)$ |
| :--- | :--- | :--- | :--- |
| BNL 1981 [60] | $1.07(6)$ | $1.07(5)$ | $1.05(5)$ |
| ANL 1982 [59] | $1.05(5)$ | $1.05(5)$ | $1.02(5)$ |
| FNAL 1983 [61] | $1.05_{-0.16}^{+0.12}$ | $1.20(11)$ | $1.17(10)$ |

Two further complications result in technical subtlety but do not affect the fit results. First, the binned event rate $d N / d E_{\nu}$ for ANL1982 is provided in a prior publication [73] that used a subset of about half the events. A second complication is the finite bin width of the $d N / d E_{\nu}$ distributions, which would yield unphysical discontinuities when displaying ANL and BNL $d N / d Q^{2}$ spectra at best fit. This effect is the result of convoluting a low energy flux with a differential cross section that has an energy-dependent kinematic limit. We use an interpolation algorithm to produce smoothed fluxes with 500 bins in energy over the original range of data. Nearly identical fit results are obtained regardless of whether the interpolation is a cubic spline, linear, or whether the original binning is used, so this step is primarily cosmetic. The smoothed and unsmoothed $E_{\nu}$ distributions are included as supplementary material to the present paper [72].

## C. Dipole fits

Our results for the axial form factor will differ from the analyses in the original publications. These differences arise from a number of sources: updated numerical inputs in Table I; not using unbinned likelihood fits; and differences in axial form factor shape assumptions. In order to understand these differences, we begin by restricting attention to the dipole ansatz,

$$
\begin{equation*}
F_{A}^{\text {dipole }}\left(q^{2}\right)=F_{A}(0)\left(1-\frac{q^{2}}{m_{A}^{2}}\right)^{-2} \tag{9}
\end{equation*}
$$

and compare to fits in the original publications.
Table II gives results for fits to the dipole ansatz (9) for the axial form factor. The table shows "fluxindependent" results from the original experiments, which performed unbinned likelihood fits to event-level data. Our results are from a Poisson likelihood fit to the binned $Q^{2}$ distribution of events obtained with a neutrino flux given by smoothing the binned reconstructed neutrino energy distribution (divided by theoretical cross section), as described in Sec. II B. Fits to the binned loglikelihood function are found by minimizing the function
$-2 \log \left[\mathcal{L}\left(F_{A}\right)\right]=2 \sum_{i}\left[\mu_{i}\left(F_{A}\right)-n_{i}+n_{i} \log \left(\frac{n_{i}}{\mu_{i}\left(F_{A}\right)}\right)\right]$,
where $n_{i}$ is the number of events in the $i$-th bin, and $\mu_{i}$ is the theory prediction (7) for the bin. Errors correspond to changes of 1.0 in the -2LL function.

Because we do not use an unbinned likelihood fit, we do not expect precise agreement even when the original choices of constants in Table I are used. Comparing the first two columns of Table II, the size of the resulting statistical uncertainties are approximately equal, and only FNAL shows a discrepancy in central value. A similar exercise was performed in Refs. [65, 74, 75], and similar results were obtained. Having reproduced the original analyses to the extent possible, we will proceed with the updated constants as in the final column of Table I.

## III. $z$ EXPANSION ANALYSIS

The dipole assumption (9) on the axial form factor shape represents an unquantified systematic error. We now remove this assumption, enforcing only the known analytic structure that the form factor inherits from QCD. We investigate the constraints from deuterium data in this more general framework. A similar analysis may be performed using future lattice QCD calculations in place of deuterium data.

## A. $z$ expansion formalism

The axial form factor obeys the dispersion relation,

$$
\begin{equation*}
F_{A}\left(q^{2}\right)=\frac{1}{\pi} \int_{t_{\mathrm{cut}}}^{\infty} d t^{\prime} \frac{\operatorname{Im} F_{A}\left(t^{\prime}+i 0\right)}{t^{\prime}-q^{2}} \tag{11}
\end{equation*}
$$

where $t_{\text {cut }}=9 m_{\pi}^{2}$ represents the leading three-pion threshold for states that can be produced by the axial current. The presence of singularities along the positive real axis implies that a simple Taylor expansion of the form factor in the variable $q^{2}$ does not converge for $\left|q^{2}\right| \geq 9 m_{\pi}^{2} \approx 0.18 \mathrm{GeV}^{2}$. Consider the new variable obtained by mapping the domain of analyticity onto the unit circle [31],

$$
\begin{equation*}
z\left(q^{2}, t_{\mathrm{cut}}, t_{0}\right)=\frac{\sqrt{t_{\mathrm{cut}}-q^{2}}-\sqrt{t_{\mathrm{cut}}-t_{0}}}{\sqrt{t_{\mathrm{cut}}-q^{2}}+\sqrt{t_{\mathrm{cut}}-t_{0}}} \tag{12}
\end{equation*}
$$

where $t_{0}$, with $-\infty<t_{0}<t_{\text {cut }}$, is an arbitrary number that may be chosen for convenience. In terms of the new variable we may write a convergent expansion,

$$
\begin{equation*}
F_{A}\left(q^{2}\right)=\sum_{k=0}^{k_{\max }} a_{k} z\left(q^{2}\right)^{k} \tag{13}
\end{equation*}
$$

where the expansion coefficients $a_{k}$ are dimensionless numbers encoding nucleon structure information.

In any given experiment, the finite range of $Q^{2}$ implies a maximal range for $|z|$ that is less than unity. We denote

TABLE III. Maximum value of $|z|$ for different $Q^{2}$ ranges and choices of $t_{0} . t_{0}^{\text {optimal }}$ is defined in Eq. (14).

| $Q_{\max }^{2}\left[\mathrm{GeV}^{2}\right]$ | $t_{0}$ | $\|z\|_{\max }$ |
| :---: | :---: | :---: |
| 1.0 | 0 | 0.44 |
| 3.0 | 0 | 0.62 |
| 1.0 | $t_{0}^{\text {optimal }}\left(1.0 \mathrm{GeV}^{2}\right)=-0.28 \mathrm{GeV}^{2}$ | 0.23 |
| 3.0 | $t_{0}^{\text {optimal }}\left(1.0 \mathrm{GeV}^{2}\right)=-0.28 \mathrm{GeV}^{2}$ | 0.45 |
| 3.0 | $t_{0}^{\text {optimal }}\left(3.0 \mathrm{GeV}^{2}\right)=-0.57 \mathrm{GeV}^{2}$ | 0.35 |

by $t_{0}^{\text {optimal }}\left(Q_{\max }^{2}\right)$ the choice which minimizes the maximum size of $|z|$ in the range $-Q_{\max }^{2} \leq q^{2} \leq 0$. Explicitly,

$$
\begin{equation*}
t_{0}^{\mathrm{optimal}}\left(Q^{2}\right)=t_{\mathrm{cut}}\left(1-\sqrt{1+Q_{\max }^{2} / t_{\mathrm{cut}}}\right) \tag{14}
\end{equation*}
$$

Table III displays $|z|_{\text {max }}$ for several choices of $Q_{\text {max }}^{2}$ and $t_{0}$.

The choice of $t_{0}$ can be optimized for various applications. We have in mind applications with data concentrated below $Q^{2}=1 \mathrm{GeV}^{2}$, and therefore take as default choice,

$$
\begin{equation*}
\bar{t}_{0}=t_{0}^{\text {optimal }}\left(1 \mathrm{GeV}^{2}\right) \approx-0.28 \mathrm{GeV}^{2} \tag{15}
\end{equation*}
$$

minimizing the number of parameters that are necessary to describe data in this region. Inspection of Table III shows that the form factor expressed as $F_{A}(z)$ becomes approximately linear. For example, taking $|z|_{\max }=0.23$ implies that quadratic, cubic, and quartic terms enter at the level of $\sim 5 \%, 1 \%$ and $0.3 \%$.

The asymptotic scaling prediction from perturbative QCD [76], $F_{A} \sim Q^{-4}$, implies the series of four sum rules [35]

$$
\begin{equation*}
\sum_{k=n}^{\infty} k(k-1) \cdots(k-n+1) a_{k}=0, \quad n=0,1,2,3 . \tag{16}
\end{equation*}
$$

We enforce the sum rules (16) on the coefficients, ensuring that the form factor falls smoothly to zero at large $Q^{2}$. Together with the $Q^{2}=0$ constraint, this leaves $N_{a}=k_{\max }-4$ free parameters in Eq. (13). From Eq. (16), it can be shown [35] that the coefficients behave as $a_{k} \sim k^{-4}$ at large $k$. We remark that the dipole ansatz (9) implies the coefficient scaling law $\left|a_{k}\right| \sim k$ at large $k$, in conflict with perturbative QCD.

In addition to the sum rules, an examination of explicit spectral functions and scattering data [31] motivates the bound of

$$
\begin{equation*}
\left|a_{k} / a_{0}\right| \leq 5 \tag{17}
\end{equation*}
$$

As noted above, from Eq. (16), the coefficients behave as $a_{k} \sim k^{-4}$ at large $k$. We invoke a fall-off of the coefficients at higher order in $k$,

$$
\begin{equation*}
\left|a_{k} / a_{0}\right| \leq 25 / k, \quad k>5 \tag{18}
\end{equation*}
$$

The bounds are enforced with a Gaussian penalty on the coefficients entering the fit. We investigate fits using a range of $k_{\max }$, other choices of $t_{0}$, and alternatives to Eqs. (17) and (18), which are briefly reported in Sec. IV.

## B. $z$ expansion basic fit results

Using the same datasets and constants as described in Sec. II and summarized in Table I, we perform fits replacing dipole axial form factor with $z$ expansion as in Eq. (13). We use the scheme choice (15), enforce the sum rule constraints (16), and use the default bounds on the coefficients $a_{k}$ in Eqs. (17), (18). The results are summarized in Table IV and displayed in Figs. 1 and 2. The coefficients corresponding to the fits with $N_{a}=4$ free parameters in Table IV are

$$
\begin{align*}
& {\left[a_{1}, a_{2}, a_{3}, a_{4}\right]} \\
& =\left\{\begin{array}{lr}
{[2.24(10), 0.6(1.0),-5.4(2.4), 2.2(2.7)]} & (\mathrm{BNL}) \\
{[2.25(10), 0.2(0.9),-4.9(2.3), 2.7(2.7)]} & (\mathrm{ANL}), \\
{[2.02(14),-1.2(1.5),-0.7(2.9), 0.1(2.8)]} & (\text { FNAL })
\end{array}\right. \tag{19}
\end{align*}
$$

where (symmetrized) errors correspond to a change of 1.0 in the -2LL function.

Table IV summarizes $z$ expansion fits with different numbers of free parameters. Focusing on the first order coefficient,

$$
\begin{align*}
& {\left[a_{1}(\mathrm{BNL}), a_{1}(\mathrm{ANL}), a_{1}(\mathrm{FNAL})\right]} \\
& = \begin{cases}{[2.23(10),} & 2.23(10), 2.02(14)], \\
{[2.24(10),} & N_{a}=3 \\
{[2.25(10), 2.02(14)],} & N_{a}=4\end{cases} \tag{20}
\end{align*}
$$

As discussed after Eq. (15), $z^{2}, z^{3}$, $z^{4}$, etc., terms in the $z$ expansion become increasingly irrelevant, corresponding to $|z|_{\max } \ll 1$ in Table III. This is borne out by the data, which determines a form factor with coefficients in Eq. (19) of order 1.0 that mostly don't push the Gaussian bounds, and a leading coefficient in Eq. (20) that is approximately the same regardless of whether terms beyond order $z^{3}$ are included.

The axial "charge" radius is defined via the form factor slope at $q^{2}=0$,

$$
\begin{equation*}
\left.\frac{1}{F_{A}(0)} \frac{d F_{A}}{d q^{2}}\right|_{q^{2}=0} \equiv \frac{1}{6} r_{A}^{2} \tag{21}
\end{equation*}
$$

For a general scheme choice $t_{0} \neq 0$, this quantity depends on all the coefficients in the $z$ expansion. Table IV illustrates that $r_{A}$ is poorly constrained without the restrictive dipole assumption. We will provide a final value for the axial radius from deuterium data after discussion of systematic errors in the next section.

The normalization factor $\mathcal{N}_{\text {fit }}$ is also included in Table IV. This parameter is allowed to float without


FIG. 1. Experimental data and best fit curves corresponding to dipole and $N_{a}=4 z$ expansion in Table IV, for BNL1981 (top pane), ANL1982 (middle pane) and FNAL1983 (bottom pane).

TABLE IV. Fits to $z$ expansion using the same data and constants as the final column of Table II. "LL" denotes log likelihood. Errors on $z$ expansion determinations of $r_{A}^{2}$ are determined from the error matrix, all others correspond to $\Delta(-2 \mathrm{LL})=1$. $N_{a}=k_{\max }-4$ denotes the number of free expansion coefficients in the $z$ expansion fit (13) with scheme choice (15), sum rule constraints (16), and bounds (17), (18). The final column is the number of bins, including bins with zero data. For $N_{a}=4$ the resulting fit parameters are displayed in Eq. (19).

|  | dipole |  |  | $N_{a}=3$ |  |  | $N_{a}=4$ |  |  | $N_{a}=5$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| experiment | -2LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | -2LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | -2LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | -2LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | $N_{\text {bins }}$ |
| BNL1981 | 70.9 | $1.14_{-0.07}^{+0.08}$ | 0.424(44) | 76.1 | $1.14_{-0.11}^{+0.12}$ | 0.36(21) | 73.4 | $1.13_{-0.11}^{+0.13}$ | 0.25(21) | 71.0 | $1.13_{-0.12}^{+0.13}$ | 0.18(21) | 49 |
| ANL1982 | 58.6 | $1.15_{-0.06}^{+0.06}$ | 0.444(44) | 62.3 | $1.15{ }_{-0.09}^{+0.10}$ | 0.38(19) | 60.9 | $1.14_{-0.10}^{+0.10}$ | 0.31(19) | 59.9 | $1.14{ }_{-0.10}^{+0.11}$ | 0.27 (19) | 49 |
| FNAL1983 | 38.2 | $1.17{ }_{-0.13}^{+0.16}$ | 0.337 (61) | 39.1 | $1.21_{-0.20}^{+0.24}$ | 0.61(28) | 39.1 | $1.21_{-0.21}^{+0.25}$ | 0.60(28) | 39.1 | $1.20_{-0.21}^{+0.26}$ | 0.58(32) | 29 |

bounds, but returns values consistent with the approximation (8) to the expectation (6).

An interesting feature of the fits displayed in Fig. 1 is that whereas the best-fit $d N / d Q^{2}$ curves for dipole and $z$ expansion are very similar in the considered $Q^{2}$ range, derived observables such as the radius in Table IV, and the absolutely normalized cross section in Fig. 2, can be markedly different. The presence of the $Q_{\min }^{2}$ cut, and the lack of an absolutely normalized flux, explains this situation, which is most apparent for FNAL1983. To illustrate, Fig. 3 shows the absolutely normalized $d \sigma^{n} / d Q^{2}$ computed using the central value dipole and $z$-expansion axial form factors for FNAL1983 in Fig. 1 and Fig. 2. Omitting the lowest- $Q^{2}$ data, and applying an overall normalization factor obscures the difference between these curves.

The normalization parameter $\mathcal{N}_{\text {fit }}$ appearing in Eq. (7) is not externally constrained in our shape fits. The uncertainty after fitting yields $\sim \pm 10 \%$ for BNL1981 and ANL1982 and $\sim \pm 20 \%$ for FNAL1983, which is significantly larger than the $\sim 3 \%$ to $\sim 5 \%$ uncertainty from Poisson statistics. A simple Poisson constraint would not be adequate considering uncertainties from acceptance and deuterium corrections described later. A rate+shape fit with a correctly motivated uncertainty on $\mathcal{N}_{\text {fit }}$ could in principle produce a somewhat better constrained form factor and cross section.

## C. Residuals analysis

The best fits are still a relatively poor description of the data, apparent in both Table IV and Fig. 1. This was briefly discussed in the thesis [77] that accompanies the ANL1982 publication: the theoretical curve is too high at very low $Q^{2}$, becoming too low above $0.2 \mathrm{GeV}^{2}$, and too high again around $0.7 \mathrm{GeV}^{2}$. Similarly, the BNL1981 publication discusses the possibility of residual scanning biases with a kinematic dependence that mimics evidence for second-class currents violating the symmetries of QCD (cf. Ref. [60], Fig. 5). These observations motivate a careful examination of systematic uncertainties assigned in the fits.

The preference of the experiments for a common $Q^{2}$ -
dependent distortion can be illustrated by comparing the residual discrepancy between the data and the best fit curves from Fig. 1 in a single plot, shown in Fig. 4. ${ }^{7}$ The distortion at lowest $Q^{2}$ is clearly significant. The data also seem to agree on potential distortions in the range $0.25<Q^{2}<3.0$. However the null hypothesis, that the data in this range were drawn from a flat distribution, yields P -value of 0.12 and is not exceptional. In order to use a $\chi^{2}$ fit for this P -value and to improve the plot readability, the upper bins in each data set were combined.

Form factors described by the $z$ expansion, hence any form factor consistent with QCD, cannot accommodate such localized distortions of the $Q^{2}$ spectrum (the dipole ansatz similarly cannot accommodate such distortions). The $R\left(Q^{2}\right)$ model for deuterium used by the original experiments asymptotically approaches unity and also does not cause such distortions. It is interesting to consider whether more complete deuteron correction models, such as Ref. [70] (considered below in Fig. 6), could produce such distortions. ${ }^{8}$ Finally, the impact of residual scanning biases should also be accounted for. In the next sections we turn to the question of assigning a systematic uncertainty to account for such effects.

## IV. SYSTEMATIC TESTS

Fits using different choices for constructing the $z$ expansion form factors should yield equivalent results for physical observables: a dependence on such choices would indicate an underestimated systematic uncertainty. Similarly, fits using different ranges of $Q^{2}$ should yield equivalent results.

[^4]

FIG. 2. Best fit curves and errors propagated from deuterium to free-neutron cross section, for BNL1981 (top pane), ANL1982 (middle pane) and FNAL1983 (bottom pane). Blue (horizontal stripes) corresponds to dipole and red (vertical stripes) to $N_{a}=4 z$ expansion in Table IV.


FIG. 3. Absolutely normalized $d \sigma^{n} / d Q^{2}$ at $E_{\nu}=10 \mathrm{GeV}$ for dipole (blue) and $z$-expansion axial form factor central values as in the FNAL1983 results of Fig. 1 and Fig. 2.


FIG. 4. Data divided by best fit prediction for the $Q^{2}$ distributions displayed in Fig. 1, for BNL(blue) ANL (red), and FNAL (green). Calculated $\chi^{2} / N_{\text {bins }}$ are $35.3 / 22,41.2 / 25$, and 10.7/14 for BNL, ANL, and FNAL respectively.

## A. Form factor scheme dependence

A test with variations of the number of free parameters was presented in Eq. (20) of the previous section. In order to translate other test fits into parameters that can be compared side-by-side, we will consider in all cases the dimensionless shape parameter defined by

$$
\begin{equation*}
\left.\bar{a}_{1} \equiv a_{1}\right|_{t_{0}=\bar{t}_{0}} \equiv-4\left(t_{\mathrm{cut}}-\bar{t}_{0}\right) F_{A}^{\prime}\left(\bar{t}_{0}\right) \tag{22}
\end{equation*}
$$

where $\bar{t}_{0} \equiv t_{0}^{\text {optimal }}\left(1 \mathrm{GeV}^{2}\right)=-0.28 \mathrm{GeV}^{2}$, as in Eq. (15). To motivate the choice (22), note that since $z$ is a small parameter, the form factor is approximately linear when expressed as a function of $z$. The slope of this approximately linear function is the essential shape parameter determined by the data, and for convenience we define the slope at $q^{2}=\bar{t}_{0}$. [The axial radius is similarly defined as the form factor slope at $q^{2}=0$ in Eq. (21).]

## 1. Magnitude of bound

Consider first the numerical value of the bound (17). For definiteness, we impose a coefficient falloff, $a_{k} \sim 1 / k$, as in Eq. (18). Focusing on $N_{a}=4$,

$$
\begin{align*}
& {\left[\bar{a}_{1}(\mathrm{BNL}), \bar{a}_{1}(\mathrm{ANL}), \bar{a}_{1}(\mathrm{FNAL})\right]} \\
& = \begin{cases}{[2.18(8), 2.17(8), 2.01(12)],} & \left|\frac{a_{k}}{a_{0}}\right| \leq \min \left(3, \frac{15}{k}\right) \\
{[2.23(10), 2.25(10), 2.02(14)],} & \left|\frac{a_{k}}{a_{0}}\right| \leq \min \left(5, \frac{25}{k}\right) \\
{[2.36(15), 2.41(15), 2.02(17)],} & \left|\frac{a_{k}}{a_{0}}\right| \leq \min \left(10, \frac{50}{k}\right)\end{cases} \tag{23}
\end{align*}
$$

Results are consistent within errors. The very conservative bound $\left|a_{k} / a_{0}\right| \leq 10$ would lead to an error that is $\sim 50 \%$ larger than our default $\left|a_{k} / a_{0}\right| \leq 5$.

## 2. Choice of $t_{0}$

Next, consider the choice of $t_{0} .{ }^{9}$ A different choice of $t_{0}$ requires more parameters to achieve the same truncation error, $\sim|z|^{N_{a}+1}$. We compare the default case of $t_{0}=$ $-0.28 \mathrm{GeV}^{2}$ and $N_{a}=4$ to the case of $t_{0}=0$ and $N_{a}=$ $7,{ }^{10}$ finding

$$
\begin{align*}
& {\left[\bar{a}_{1}(\mathrm{BNL}), \bar{a}_{1}(\mathrm{ANL}), \bar{a}_{1}(\mathrm{FNAL})\right]} \\
& = \begin{cases}{[2.24(10), 2.25(10), 2.02(14)]} & \left(N_{a}=4, t_{0}=\bar{t}_{0}\right) \\
{[2.22(9), 2.21(10), 2.02(14)]} & \left(N_{a}=7, t_{0}=0\right)\end{cases} \tag{24}
\end{align*}
$$

where the errors are propagated using the covariance matrix for the coefficients $a_{k}$. Nearly identical results are obtained for different choices of $t_{0}$.

[^5]
## B. Subsets of the $Q^{2}$ range

A nonstatistical scatter of data points about the best fit curves is apparent in Fig. 1, and indicated by the poor fit quality in Table IV. Removing subsets of the data at high or low $Q^{2}$ will help isolate sources of tension between data and fit.

First, consider the removal of high $Q^{2}$ data, fitting to bins whose center is within the restricted range $Q^{2} \leq$ $1 \mathrm{GeV}^{2}$. The analog of Table IV for this case is given by Table V. Figure 5 shows comparisons of best fit curves and data points. The analog of Eq.(19) is

$$
\begin{align*}
& {\left.\left[a_{1}, a_{2}, a_{3}, a_{4}\right]\right|_{Q^{2} \leq 1 \mathrm{GeV}^{2}}} \\
& =\left\{\begin{array}{lr}
{[1.99(15), 0.5(1.1),-3.6(2.6), 1.1(2.7)]} & (\mathrm{BNL}) \\
{[2.29(14), 0.2(0.9),-5.2(2.5), 2.9(2.7)]} & (\mathrm{ANL}) . \\
{[1.88(25),-0.9(1.6),-0.3(2.9),-0.3(2.8)]} & (\mathrm{FNAL})
\end{array}\right. \tag{25}
\end{align*}
$$

The omission of low- $Q^{2}$ data has a similarly large effect on the fit parameters. Fitting to the range $Q^{2} \geq$ $0.2 \mathrm{GeV}^{2}$, the results are given in Table VI. The $z$ expansion coefficients are determined for $N_{a}=4$ to be

$$
\begin{align*}
& {\left.\left[a_{1}, a_{2}, a_{3}, a_{4}\right]\right|_{Q^{2} \geq 0.2 \mathrm{GeV}^{2}}} \\
& = \begin{cases}{[2.35(10),-2.0(1.2),-1.4(2.8), 1.4(2.7)]} & (\text { BNL }) \\
{[2.34(10),-3.6(1.2), 1.6(2.8), 0.9(2.8)]} & (\text { ANL }) \\
{[2.04(16),-1.3(1.6),-0.5(3.0), 0.1(2.8)]} & (\text { FNAL ) }\end{cases} \tag{26}
\end{align*}
$$

Comparing the results in Tables IV, V, and VI and in Eqs. (19), (25), and (26), we see that the leading $a_{1}$ and $a_{2}$ parameters shift in some cases by about twice the statistical uncertainty of the fits. This reflects how different parts of the $Q^{2}$ range contribute to tensions in the fit. The minimum value of $-2 \mathrm{LL} \sim \chi^{2}$ decreases in both cases, closer to a range that would be considered an adequate description of the data. The improvement when eliminating the low- $Q^{2}$ region is especially striking considering it amounts to only two or three bins of data in each data set.

One method to translate the tensions in the fit to an uncertainty on the fit parameters is to consider what additional error is necessary to obtain a reduced $\chi^{2}$ of unity. We include an error for each data point proportional to the number of events in the original $d N / d Q^{2}$ distribution. This requires the use of a $\chi^{2}$ calculation instead of a log-likelihood fit, which we achieve by limiting the test to the sample with $Q^{2} \leq 1 \mathrm{GeV}^{2}$. Adding this error in quadrature to the statistical error, we see that for BNL, an additional $10 \%$ error is required, while ANL requires an additional $7.5 \%$ error.

TABLE V. Same as Table IV, but fitting only to data with $Q^{2} \leq 1 \mathrm{GeV}^{2}$. For $N_{a}=4$ the resulting fit parameters are displayed in Eq. (25).

|  | dipole |  |  |  | $N_{a}=3$ |  |  |  | $N_{a}=4$ |  |  | $N_{a}=5$ |  |  |
| ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| experiment | -2 LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | -2 LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | -2 LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | -2 LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | $N_{\text {bins }}$ |  |
| BNL1981 | 24.7 | $1.16_{-0.08}^{+0.08}$ | $0.348(48)$ | 27.2 | $1.17_{-0.13}^{+0.14}$ | $0.32(22)$ | 27.0 | $1.17_{-0.13}^{+0.14}$ | $0.28(22)$ | 26.6 | $1.16_{-0.13}^{+0.14}$ | $0.24(22)$ | 16 |  |
| ANL1982 | 28.2 | $1.14_{-0.06}^{+0.07}$ | $0.452(52)$ | 31.7 | $1.15_{-0.09}^{+0.10}$ | $0.38(19)$ | 30.5 | $1.14_{-0.10}^{+0.10}$ | $0.31(20)$ | 29.2 | $1.13_{-0.11}^{+0.11}$ | $0.24(20)$ | 19 |  |
| FNAL1983 | 8.3 | $1.16_{-0.18}^{+0.26}$ | $0.33(12)$ | 8.3 | $1.22_{-0.23}^{+0.29}$ | $0.54(31)$ | 8.2 | $1.23_{-0.24}^{+0.29}$ | $0.56(29)$ | 8.1 | $1.24_{-0.24}^{+0.30}$ | $0.57(26)$ | 9 |  |

TABLE VI. Same as Table IV, but fitting only to data with $Q^{2} \geq 0.2 \mathrm{GeV}^{2}$. For $N_{a}=4$ the resulting fit parameters are displayed in Eq. (26).

|  | dipole |  |  | $N_{a}=3$ |  |  | $N_{a}=4$ |  |  | $N_{a}=5$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| experiment | -2LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | -2LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | $-2 \mathrm{LL}$ | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | -2LL | $\mathcal{N}_{\text {fit }}$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ | $N_{\text {bins }}$ |
| BNL1981 | 60.7 | $1.25_{-0.14}^{+0.21}$ | 0.61(13) | 62.4 | $1.288_{-0.17}^{+0.20}$ | 0.83(24) | 61.5 | $1.26_{-0.18}^{+0.21}$ | 0.74(25) | 60.9 | $1.25_{-0.19}^{+0.23}$ | 0.67(24) | 47 |
| ANL1982 | 43.2 | $1.400_{-0.38}^{+0.25}$ | $1.45_{-0.49}^{+0.92}$ | 45.8 | $1.32_{-0.18}^{+0.21}$ | 1.04(24) | 45.8 | $1.32_{-0.20}^{+0.23}$ | 1.03(25) | 45.8 | $1.32_{-0.21}^{+0.25}$ | 1.05(24) | 46 |
| FNAL1983 | 38.2 | $1.16{ }_{-0.16}^{+0.22}$ | 0.33(7) | 39.1 | $1.22_{-0.25}^{+0.31}$ | 0.64(31) | 39.1 | $1.22_{-0.25}^{+0.32}$ | 0.63(30) | 39.0 | $1.21_{-0.26}^{+0.34}$ | 0.60(35) | 28 |

## V. SYSTEMATIC ERRORS

The experimental uncertainties in the fits summarized in Table IV correspond only to statistical errors on the number of events in each bin. With a framework in place to quantify theoretical form factor shape uncertainty, let us examine several sources of systematic error, and their impact on the extraction of $F_{A}$.

Experimental systematic uncertainties come from the construction of the neutrino flux, and from acceptance corrections. A theoretical systematic error arises from uncertain modeling of deuteron effects.

## A. Flux

Our procedure includes a self-consistent determination of the neutrino flux for fits to the $Q^{2}$ distributions, as described in Sec. II B. Systematic uncertainty estimates in the experimenter's ab-initio flux do not apply. Instead we check for sensitivity to fluctuations in the number of events by varying one $d N / d E_{\nu}$ bin by its statistical error, reextracting fit parameters, and then repeating for all bins. Adding errors in quadrature, the result for the BNL dataset is

$$
\begin{equation*}
\bar{a}_{1}=2.24 \pm 0.10_{\text {stat } . Q^{2}} \pm 0.04_{\text {stat } . E_{\nu}} \quad(\text { BNL1981 }) \tag{27}
\end{equation*}
$$

Such an additional flux error is numerically subleading compared to statistical error, and also to the systematic error assigned below to account for deuteron and acceptance corrections. We neglect it in our final fits.

The consistency of the flux procedure could also be impacted by distortions of the $d N / d E_{\nu}$ distribution by $Q_{\text {min }}^{2}$ cuts or deuteron corrections. Recall that the energy distribution from BNL1981, but not from ANL1982 or FNAL1983, was corrected for these effects. We have
checked that the resulting variations are even smaller than the statistical fluctuations in Eq. (27), and are neglected.

## B. Acceptance corrections

One source of uncertainty, especially in the limit of very low $Q^{2}$, is the acceptance corrections associated with human-eye scanning of the bubble chamber photographs. For example, Fig. 1 of ANL 1982 [59] provides an estimate of the scanning efficiency ranging from $e=90 \pm 7 \%$ at $0.05 \mathrm{GeV}^{2}<Q^{2}<0.1 \mathrm{GeV}^{2}$ to $e=98 \pm 1 \%$ for $Q^{2}>0.15 \mathrm{GeV}^{2}$. We include a possible correlated efficiency correction by making the following replacement in the efficiency-corrected number of events:

$$
\begin{equation*}
\frac{d N}{e\left(Q^{2}\right)} \rightarrow \frac{d N}{e\left(Q^{2}\right)+\eta d e\left(Q^{2}\right)}=\frac{d N}{e\left(Q^{2}\right)}\left(1+\eta \frac{d e\left(Q^{2}\right)}{e\left(Q^{2}\right)}\right)^{-1} \tag{28}
\end{equation*}
$$

Here $\eta=0 \pm 1$ is a parameter introduced in the fit, and we use a simple linear interpolation of the function in Ref. [59] for the efficiency $e\left(Q^{2}\right)$ and efficiency error $d e\left(Q^{2}\right)$.

In the BNL dataset, an efficiency effect with similar magnitude is presented, but not directly in the $Q^{2}$ variable. For simplicity we take the ANL function to represent possible effects also in the BNL and FNAL datasets, with independent floating scale parameters $\eta=0 \pm 1$ in Eq. (28). The shape parameters and minimum -2LL values are as follows, comparing results with and without the acceptance correction,

$$
\text { BNL: } \quad\left[\bar{a}_{1},-2 \mathrm{LL}\right]= \begin{cases}{[1.99(15), 27.0]} & \text { (without) } \\ {[2.04(15), 26.0]} & \text { (with) }\end{cases}
$$





FIG. 5. Same as Fig. 1, but with $Q^{2} \leq 1 \mathrm{GeV}^{2}$. These fits correspond to the $N_{a}=4 z$ expansion in Table V.


FIG. 6. Differential scattering cross sections for neutrinodeuteron scattering at 1 GeV neutrino energy, employing different nuclear models. The solid (red) curve is the freeneutron result. The dashed (blue) curve is obtained from the free-neutron result using the model from Ref. [69], as in the original deuterium analyses. The top dot-dashed (black) curve is extracted at $E_{\nu}=1 \mathrm{GeV}$ from Ref. [70]. The charged lepton mass is neglected in this plot.

ANL: $\quad\left[\bar{a}_{1},-2 \mathrm{LL}\right]=\left\{\begin{array}{ll}{[2.29(14), 30.5]} & \text { (without) } \\ {[2.38(14), 26.3]} & \text { (with) }\end{array}\right.$,
FNAL : $\quad\left[\bar{a}_{1},-2 \mathrm{LL}\right]=\left\{\begin{array}{ll}{[1.88(25), 8.2]} & \text { (without) } \\ {[1.88(25), 8.2]} & \text { (with) }\end{array}\right.$.

The parameter $\eta$ takes on values of $-1.9,-1.0$, and +0.01 for data from ANL1982, BNL1981, and FNAL1983 respectively; the negative values indicate a pull to decrease the predicted cross section to match the data. In each case there is only modest improvement in the fit quality, and small impact on the form factor shape. Acceptance corrections within the quoted range have only minor impact.

## C. Deuteron corrections

The analysis to this point, like the original analyses, used the deuteron correction model $R\left(Q^{2}\right)$ of Singh [69]. This model yields a suppression of the cross section for $Q^{2}<0.16 \mathrm{GeV}^{2} .{ }^{11}$ An example of a modern calculation

[^6]with extended range in energy and $Q^{2}$ is given by Shen et al. in Ref. [70]. ${ }^{12}$ The Shen et al. model is overlaid with the original Singh model as well as the free neutron model in Fig. 6. The Shen et al. model deviates substantially from the free-neutron result at the $\sim 20 \%$ level over a broad $Q^{2}$ range. These models do not constitute an estimate of the uncertainty on deuteron corrections, but suggest an avenue for future work even if there are no future measurements on deuterium.

Assuming an energy independent, but $Q^{2}$ dependent, deuteron correction, the change in the fit results can be compared. For illustration, we employ the results of Ref. [70] at $E_{\nu}=1 \mathrm{GeV}$, and limit attention to $Q^{2} \leq 1 \mathrm{GeV}^{2}$, i.e., the configuration of Table V and Eq. (25). Shape parameter and minimum $-2 L L$ values are

BNL: $\quad\left[\bar{a}_{1},-2 \mathrm{LL}\right]=\left\{\begin{array}{ll}{[1.99(15), 27.0]} & (\text { Singh }) \\ {[2.16(14), 25.1]} & \text { (Shen et al.) }\end{array}\right.$,
ANL: $\quad\left[\bar{a}_{1},-2 \mathrm{LL}\right]=\left\{\begin{array}{ll}{[2.29(14), 30.5]} & (\text { Singh }) \\ {[2.46(13), 29.2]} & \text { (Shen et al.) }\end{array}\right.$,
FNAL $: \quad\left[\bar{a}_{1},-2 \mathrm{LL}\right]=\left\{\begin{array}{ll}{[1.88(25), 8.2]} & \text { (Singh) } \\ {[2.00(25), 9.1]} & \text { (Shen et al.) }\end{array}\right.$.

The extracted form factor shifts to mimic the difference in the curves in Fig. 6, and there is slight improvement in fit quality for two of the three data sets.

## D. Final systematic error budget

The most important systematic uncertainties are the two that significantly modify the $Q^{2}$ distribution: acceptance corrections and the deuteron correction. In our final analysis, we modify the original fits displayed in Table V. First, we allow a correlated acceptance correction as in Eq. (28). Second, we include a $10 \%$ error added in quadrature to statistical error in each $Q^{2}$ bin to account for residual deuteron or other systematic corrections, as described at the end of Sec. IV B. With these corrections in place, we perform a $\chi^{2}$ fit to all data up to $Q^{2}=1 \mathrm{GeV}^{2}$. The neglect of data above $Q^{2}=1 \mathrm{GeV}^{2}$ has only minor impact on the extraction of $F_{A}\left(q^{2}\right)$, and allows a simple treatment of these combined uncertainties with full covariance using a $\chi^{2}$ fit.

As an alternative, we also provide a log-likelihood fit to the data up to $Q^{2}=3 \mathrm{GeV}^{2}$, but without inflated errors to account for deuterium and other residual systematics. This has the benefit of including data over the entire kinematic range, but omits sources of systematic error that would need to be treated separately.

[^7]
## VI. AXIAL FORM FACTOR EXTRACTION

The best axial form factor is extracted from a joint fit to the three datasets. We choose $N_{a}=4$ free parameters with $t_{0}=t_{0}^{\text {optimal }}\left(1 \mathrm{GeV}^{2}\right)$ and data with $Q^{2} \leq 1 \mathrm{GeV}^{2}$. As discussed above, this corresponds to a $k_{\max }=8 z$ expansion, where five linear combinations of coefficients are fixed by the $Q^{2}=0$ constraint and by the four sum rules (16). The acceptance correction free parameter is independent for each experiment in the joint fit.

Our knowledge of the axial form factor resulting from deuterium scattering data is summarized by constraints on the coefficients $a_{k}$. Central values and $1 \sigma$ errors determined from $\Delta \chi^{2}=1$ are ${ }^{13}$
$\left[a_{1}, a_{2}, a_{3}, a_{4}\right]=[2.30(13),-0.6(1.0),-3.8(2.5), 2.3(2.7)]$.

The diagonal entries of the error (covariance) matrix, computed from the inverse of the Hessian matrix for $\chi^{2}\left(\left\{a_{k}\right\}\right)$, are

$$
\begin{equation*}
E_{\text {diag. }}=[0.0154,1.08,6.54,7.40] . \tag{32}
\end{equation*}
$$

Note that $\left(E_{\text {diag. }}\right)_{i} \approx\left(\delta a_{i}\right)^{2}$, reflecting approximately Gaussian behavior. The four-dimensional correlation matrix is

$$
C_{i j}=\left(\begin{array}{cccc}
1 & 0.350 & -0.678 & 0.611  \tag{33}\\
0.350 & 1 & -0.898 & 0.367 \\
-0.678 & -0.898 & 1 & -0.685 \\
0.611 & 0.367 & -0.685 & 1
\end{array}\right)
$$

and as usual the error matrix is given by $E_{i j}=\delta a_{i} \delta a_{j} C_{i j}$. This description can be systematically improved when and if further data or externally constrained deuterium models become available. The form factor is plotted ver$\operatorname{sus} Q^{2}$ and versus $z$ in Fig. 7, and compared with a previous world average dipole form factor from Ref. [55].

We also provide an alternate log-likelihood determination of the axial form factor to the range $Q^{2}<3.0 \mathrm{GeV}^{2}$, but without deuteron systematic corrections. Central values and $1 \sigma$ errors determined from $\Delta(-2 L L)=1$ are

$$
\begin{equation*}
\left[a_{1}, a_{2}, a_{3}, a_{4}\right]=[2.28(8), 0.25(95),-5.2(2.3), 2.6(2.7)] \tag{34}
\end{equation*}
$$

The diagonal entries of the error matrix are

$$
\begin{equation*}
E_{\text {diag }}=[0.00635,0.781,4.49,6.87] \tag{35}
\end{equation*}
$$

[^8]

FIG. 7. Final form factor from Eqs. (31), (32) and (33). Also shown is the dipole axial form factor with axial mass $m_{A}=$ 1.014(14) GeV [55].
and the four-dimensional correlation matrix is

$$
C_{i j}=\left(\begin{array}{cccc}
1 & 0.321 & -0.677 & 0.761  \tag{36}\\
0.321 & 1 & -0.889 & 0.313 \\
-0.677 & -0.889 & 1 & -0.689 \\
0.761 & 0.313 & -0.689 & 1
\end{array}\right)
$$

## VII. APPLICATIONS

Having presented the axial form factor with errors and correlations amongst the coefficients, we may systematically compute derived observables that depend on this function. We consider several applications of our results.

TABLE VII. Axial radius extracted using best values from Table I, and default priors as discussed in the text. Note that the joint fit is not an average, but a simultaneous fit to all of the datasets.

| dataset | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ <br> $\left(N_{a}=3\right)$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ <br> $\left(N_{a}=4\right)$ | $r_{A}^{2}\left[\mathrm{fm}^{2}\right]$ <br> $\left(N_{a}=5\right)$ |
| :--- | :--- | :--- | :--- |
| BNL 1981 | $0.56(23)$ | $0.52(25)$ | $0.48(26)$ |
| ANL 1982 | $0.69(21)$ | $0.63(23)$ | $0.57(24)$ |
| FNAL 1983 | $0.63(34)$ | $0.64(35)$ | $0.64(35)$ |
| Joint Fit | $0.54(20)$ | $0.46(22)$ | $0.39(23)$ |

## A. Axial radius

We begin with the axial radius, defined in Eq. (21). While the radius by itself is not the only quantity of interest to neutrino scattering observables, it is only through the $q^{2} \rightarrow 0$ limit that a robust comparison can be made to other processes such as pion electroproduction.

The form factor coefficients and error matrix from the $\chi^{2}$ fit in Sec. VI determine the radius as

$$
\begin{equation*}
r_{A}^{2}=0.46(22) \mathrm{fm}^{2} \tag{37}
\end{equation*}
$$

The constraint is much looser than would be obtained by restricting to the dipole model, cf. Table IV. ${ }^{14}$ For comparison, let us consider the constraints from individual experiments. Table VII gives results for $N_{a}=3,4,5$ free parameters, with errors determined from the error matrix in Eqs. (32) and (33). The results from individual experiments are consistent with the joint fit. Note that the joint fit is not simply the average of the individual fits. This situation arises from a slight tension between data and Gaussian coefficient constraints (17) when comparing a single data set to the statistically more powerful combined data.

## B. Neutrino-nucleon quasielastic cross sections

Current and future neutrino oscillation experiments will precisely measure neutrino mixing parameters, determine the neutrino mass hierarchy, and search for possible CP violation and other new phenomena. This program relies on accurate predictions, with quantifiable uncertainties, for neutrino interaction cross sections. As the simplest examples, consider the charged-current quasielastic cross section $\sigma\left(E_{\nu}\right)$ for neutrino (antineutrino) scattering on an isolated neutron (proton).

The best fit cross section and uncertainty are shown in Fig. 8, and compared to the prediction of dipole $F_{A}$ with axial mass $m_{A}=1.014(14)$ [55]. At representative

[^9]

FIG. 8. Free nucleon CCQE cross section computed from Eqs. (31), (32) and (33), for neutrino-neutron (top) and antineutrino-proton (bottom) scattering. Also shown are results using dipole axial form factor with axial mass $m_{A}=$ $1.014(14) \mathrm{GeV}$ [55].
energies, the cross sections and uncertainties shown in Fig. 8 are

$$
\begin{align*}
& \sigma_{\nu n \rightarrow \mu p}\left(E_{\nu}=1 \mathrm{GeV}\right)=10.1(0.9) \times 10^{-39} \mathrm{~cm}^{2} \\
& \sigma_{\nu n \rightarrow \mu p}\left(E_{\nu}=3 \mathrm{GeV}\right)=9.6(0.9) \times 10^{-39} \mathrm{~cm}^{2} \tag{38}
\end{align*}
$$

for neutrinos and

$$
\begin{align*}
& \sigma_{\bar{\nu} p \rightarrow \mu n}\left(E_{\nu}=1 \mathrm{GeV}\right)=3.83(23) \times 10^{-39} \mathrm{~cm}^{2} \\
& \sigma_{\bar{\nu} p \rightarrow \mu n}\left(E_{\nu}=3 \mathrm{GeV}\right)=6.47(47) \times 10^{-39} \mathrm{~cm}^{2} \tag{39}
\end{align*}
$$

for antineutrinos.


FIG. 9. Cross section for charged-current quasielastic events from the MINERvA experiment [56] as a function of reconstructed $Q^{2}$, compared with prediction using relativistic Fermi gas (RFG) nuclear model with $z$ expansion axial form factor extracted from deuterium data. MINERvA data uses an updated flux prediction from [83]. Also shown are results using the same nuclear model but dipole form factor with axial mass $m_{A}=1.014(14) \mathrm{GeV}$ [55].

## C. Neutrino nucleus cross sections

Connecting nucleon-level information to experimentally observed neutrino-nucleus scattering cross sections requires data-driven modeling of nuclear effects. Our description of the axial form factor and uncertainty in Eqs. (31), (32), and (33) can be readily implemented in neutrino event generators that interface with nuclear models. ${ }^{15}$

A multitude of studies and comparisons are possible. As illustration, consider MINERvA quasielastic data on carbon [56]. Figure 9 shows a comparison of the $Q^{2}$ distribution of measured events with the predictions from our $F_{A}\left(q^{2}\right)$, using a relativistic Fermi gas nuclear model in the default configuration of the GENIE v2.8 neutrino event generator [6]. For comparison, we display the result obtained using a dipole $F_{A}$ with axial mass central value and error as quoted in the world average of Ref. [55]. The central curves differ in their kinematic dependence, and the dipole result severely underestimates the uncertainty propagated from deuterium data.

The $z$ expansion implementation within GENIE in-

[^10]cludes a complete description of parameter errors and correlations. This will provide a systematic approach for testing different nuclear models and fitting nuclear model parameters, and for propagating uncertainties in nucleonlevel amplitudes through to oscillation observables.

## D. Discussion

The dipole ansatz has been commonly used to parameterize the axial form factor in neutrino cross section predictions. The axial mass parameter in this ansatz often appears with either a very small uncertainty, e.g. $m_{A}=1.014(14) \mathrm{GeV}$ [55], or a very large uncertainty, e.g. $m_{A}=1.21(45) \mathrm{GeV}[14]$.

In the first case, the small error estimate results from the restrictive dipole ansatz, and is likely an underestimate of the actual uncertainty: as a point of comparison, the $\lesssim 1.5 \%$ axial radius error is comparable to or smaller than the uncertainty on the proton charge radius $[35,68]$. Recall that the charge radius is defined for the vector charge form factor analogously to the axial radius for the axial form factor. In contrast to the axial radius from neutrino-deuteron scattering, the charge radius from electron-proton scattering involves much higher statistics, a monoenergetic beam, and a simpler, proton, target.

In the second case, the large uncertainty on $m_{A}$ is typically included to account for tensions in external inputs from other experiments [14], and/or poorly constrained nuclear effects. Neither of these approaches is suited to the kinds of analyses that can be undertaken with modern cross section data such as the MINERvA example considered in Fig. 9. Underestimating nucleon-level uncertainties will bias conclusions about neutrino parameters or nuclear models. Inflating errors on $m_{A}$ within a dipole ansatz fails to capture the correct kinematic dependence of either nucleon-level uncertainties, or of nuclear corrections ${ }^{16}$.

## VIII. SUMMARY AND CONCLUSION

The constraints of elementary target data are critical to precision neutrino-nucleus cross sections underlying the accelerator neutrino program. Oscillation experiments rely on event rate predictions using nucleon-level amplitudes corrected for nuclear effects. Cross section experiments on nuclear targets can measure these nuclear effects but a complete accounting of uncertainty
in nucleon-level amplitudes is critical for disentangling nucleon-level, nuclear-level, and flux uncertainties, and for determining final sensitivity to fundamental neutrino parameters.

The axial form factor is a prominent source of nucleonlevel uncertainty. We have analyzed the world dataset for quasielastic neutrino-deuteron scattering using a modelindependent description of the axial form factor. Our final results are presented with central values, (31), errors (32) and correlations (33). Any observable depending on the axial form factor may be computed from these results, with a complete error budget.

The axial radius, governing the shape of the axial form factor, is presented in Eq. (37). It has a significantly larger uncertainty than previously estimated based on the unjustified dipole ansatz. Benchmark total cross sections on nucleon targets are presented in Fig. 8 and Eqs. (38), (39). The incorporation of nuclear effects with the RFG model is illustrated in Fig. 9.

The form factor and uncertainty budget presented here are important new inputs to the neutrino cross section effort. It is interesting to investigate potential impacts and interplay with a variety of other processes such as neutrinoless double beta decay matrix elements [84, 85] and the muon capture rate in muonic hydrogen [30]. The methodology presented can be revised or extended if new information becomes available. Future hydrogen or deuterium data would be trivial to include. Updated calculations for neutrino-deuteron scattering, especially if accompanied by an uncertainty, can be readily incorporated on top of this result. Lattice QCD holds promise to determine the axial form factor over much of the relevant $Q^{2}$ range, in a manner that is free from nuclear corrections [86-91].

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[^1]:    ${ }^{1}$ Formalism for $z$ expansion and nucleon form factors is described in Refs. [31, 32], and several applications are found in Refs. [3336]. Related formalism and applications may be found in [37-53].
    2 An updated BNL dataset was presented in Ref. [62] with a factor $\approx 2$ increase in number of events. However, we were unable to extract a sufficiently precise $Q^{2}$ distribution of events from this reference, since the data were presented on a logarithmic scale

[^2]:    ${ }^{5}$ For BNL1981 and ANL1982, the digitized number of events in each $Q^{2}$ bin was rounded to the nearest integer, resulting in the same total numbers, 1236 and 1792 respectively, quoted in the original publications. For FNAL1983, the digitization produced near-integer results in each $Q^{2}$ bin, but the total summed event number, 354 , differs from the value 362 quoted in the original publication.

[^3]:    ${ }^{6}$ While $E_{\nu}=1 \mathrm{GeV}$ is close to the peak energy for the BNL1981 and ANL1982 datasets, the FNAL1983 dataset involved higher energy. However, these ratios have mild energy dependence above $E_{\nu} \sim 1 \mathrm{GeV}$, e.g. at $E_{\nu}=10 \mathrm{GeV}$ the result is $\hat{\mathcal{N}}\left(Q_{\text {min }}^{2}=\right.$ $0.10 \mathrm{GeV}^{2}$ ) $\approx 1.25$.

[^4]:    ${ }^{7}$ For definiteness, the best fit curve is from a simultaneous fit to the BNL, ANL and FNAL datasets. A nearly identical plot is obtained if different best fit curves for each dataset are used.
    ${ }^{8}$ Calculations of multi-nucleon effects for heavier nuclei like carbon exhibit qualitatively similar characteristics throughout this region of $Q^{2}[78,79]$.

[^5]:    ${ }^{9}$ For $t_{0}=t_{0}^{\text {optimal }}\left(1 \mathrm{GeV}^{2}\right)=-0.28 \mathrm{GeV}^{2}$, by design, the shape parameter is identified with the linear coefficient of the $z$ expansion in Eq. (13). Since $\bar{a}_{1}$ (Eq. 22) is a physical observable, it can be computed for any choice of $t_{0} \neq \bar{t}_{0}$.
    10 Both cases have $|z|_{\max }^{N_{a}+1} \approx 0.02$ in the range $0<Q^{2}<3 \mathrm{GeV}^{2}$.

[^6]:    11 A follow-up analysis [80] considers effects of meson exchange currents and alternate deuteron wave functions, with a total result very similar to Ref. [69].

[^7]:    12 See also Ref. [81].

[^8]:    13 The complete specification for the form factor involves the normalization $g_{A}=-1.2723$ from Table I; the pion mass $m_{\pi}=$ 0.14 GeV employed in the specification of $t_{\text {cut }}=9 m_{\pi}^{2}$ in Eq. (12); and the choice $t_{0}=-0.28 \mathrm{GeV}^{2}$. The remaining coefficients, $a_{0}, a_{5}, a_{6}, a_{7}$ and $a_{8}$, are determined by $F_{A}(0)=g_{A}$, and by the sum rule constraints (16); for ease of comparison we list the complete list of central values here: $\left[a_{0}, \cdots, a_{8}\right]=$ $[-0.759,2.30,-0.6,-3.8,2.3,2.16,-0.896,-1.58,0.823]$.

[^9]:    ${ }^{14}$ Extractions of the radius from electroproduction data are also strongly influenced by the dipole assumption [31].

[^10]:    15 The $z$ expansion will be available in GENIE production release v2.12.0. The code is currently available in the GENIE trunk prior to its official release. The module provides full generality of the $z$ expansion, and supports reweighting and error analysis with correlated parameters.

[^11]:    16 Non-dipole parameterizations have been considered in [66, 82].

